

System Identification

Lecture 6

PE Method - Linear Regression and Approximate Modelling

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Linear regression

What are the attractive properties of model structures that are linear-in-the-parameters?

Linear regression

Particular model structures: ARX, FIR

Prediction:

$$\begin{aligned}\hat{y}(t|t-1) &= H^{-1}Gu(t) + [1 - H^{-1}]y(t) \\ &= B(q^{-1})u(t) + [1 - A(q^{-1})]y(t)\end{aligned}$$

Consequence:

$$\hat{y}(t|t-1; \theta) = \varphi^T(t)\theta$$

with

$$\begin{aligned}\varphi(t) &= [-y(t-1) \cdots -y(t-n_a) \ u(t) \cdots u(t-n_b+1)]^T \\ \theta &= [a_1 \ a_2 \ \cdots \ a_{n_a} \ b_0 \ \cdots \ b_{n_b-1}]^T\end{aligned}$$

Predictor is **linear-in-the-parameters** θ .

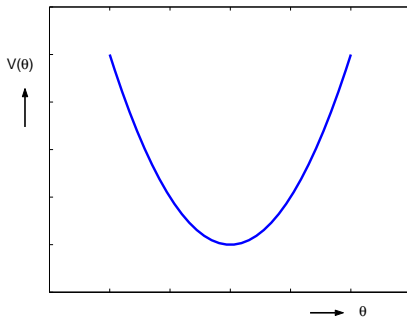
Criterion:

$$V_N(\theta) = \frac{1}{N} \sum_{t=0}^{N-1} \varepsilon(t, \theta)^2$$

with

$$\varepsilon(t, \theta) = y(t) - \varphi^T(t)\theta$$

is **quadratic** in θ .



(for scalar θ)

Minimum is uniquely achieved in $\hat{\theta}_N$ by solving

$$\left. \frac{\partial V_N(\theta)}{\partial \theta} \right|_{\theta=\hat{\theta}_N} = 0$$

$$\begin{aligned}
 V_N(\theta, Z^N) &= \frac{1}{N} \sum_{t=1}^N [y(t) - \varphi^T(t)\theta]^2 \\
 \frac{\partial V_N(\theta, Z^N)}{\partial \theta} &= 2 \frac{1}{N} \sum_{t=1}^N [y(t) - \varphi^T(t)\theta] \frac{-\partial \varphi^T(t)\theta}{\partial \theta} \\
 &= -2 \frac{1}{N} \sum_{t=1}^N \underbrace{[y(t) - \varphi^T(t)\theta]}_{\text{scalar}} \cdot \underbrace{\varphi(t)}_{\text{vector}} \\
 &= -2 \frac{1}{N} \sum_{t=1}^N [\varphi(t)y(t) - \varphi(t)\varphi^T(t)\theta]
 \end{aligned}$$

Setting derivative to 0 in $\theta = \hat{\theta}_N$ delivers:

$$\left[\frac{1}{N} \sum_{t=1}^N \varphi(t)\varphi^T(t) \right] \hat{\theta}_N = \frac{1}{N} \sum_{t=1}^N \varphi(t)y(t)$$

As a consequence:

$$\hat{\theta}_N = \left[\underbrace{\frac{1}{N} \sum_{t=1}^N \varphi(t) \varphi^T(t)}_{R(N)} \right]^{-1} \cdot \underbrace{\frac{1}{N} \sum_{t=1}^N \varphi(t) y(t)}_{f(N)}$$

- Analytical solution through simple matrix operations.
- All terms on the right hand side are provided by the data
- Requirement for unique solution: $R(N)$ is invertible.
This comes down to a sufficient excitation of the input signal.

Analysis:

If the data generating system is given by

$$y(t) = \varphi^T(t)\theta_0 + w_0(t)$$

which is equivalent to

$$y(t) = \frac{B_0(q^{-1})}{A_0(q^{-1})}u(t) + \frac{1}{A_0(q^{-1})}w_0(t)$$

Then

$$\hat{\theta}_N = \theta_0 + R(N)^{-1} \cdot \frac{1}{N} \sum_{t=1}^N \varphi(t)w_0(t).$$

The second (bias) term disappears if $w_0(t)$ uncorrelated to the components in $\varphi(t)$.

$$\bar{\mathbb{E}}\{\varphi(t)w_0(t)\} = \bar{\mathbb{E}} \begin{bmatrix} -y(t-1) \\ \vdots \\ -y(t-n_a) \\ u(t) \\ \vdots \\ u(t-n_b+1) \end{bmatrix} w_0(t)$$

For the asymptotic bias to be 0, $w_0(t)$ has to satisfy

$$\bar{\mathbb{E}}\{\varphi(t)w_0(t)\} = 0.$$

For ARX models (y - and u -part of regressor are present):
Asymptotically unbiased if w_0 is white noise.

This condition is satisfied if $\mathcal{S} \in \mathcal{M}$.

For FIR models (y -part of regressor is not present):
Asymptotically unbiased if w_0 uncorrelated to u

This condition is satisfied if $G_0 \in \mathcal{G}$.

Note the relation with the consistency results of Lecture 5:

- ▶ In an ARX model structure G_0 can be estimated consistently if $\mathcal{S} \in \mathcal{M}$;
- ▶ In an FIR model structure G_0 can be estimated consistently if $G_0 \in \mathcal{G}$
(because of the independent parametrization of G and H)

Summary

Model structures (ARX,FIR) leading to linear regression problems have become very popular, because of their computational simplicity.

Approximate Modelling

What can be said about

$$G(q, \theta^*) \quad H(q, \theta^*)$$

if $S \notin \mathcal{M}$, and even $G_0 \notin \mathcal{G}$?

Can we characterize the approximative properties of identified models if they can not capture all dynamics of the data generating system?

We know (convergence result) that:

$$\theta^* = \arg \min_{\theta} \bar{V}(\theta)$$

and

$$\begin{aligned} \bar{V}(\theta) &= \bar{\mathbb{E}}_{\mathcal{E}}(t, \theta)^2 \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\mathcal{E}}(\omega) d\omega \end{aligned}$$

(Parseval; both expressions are equal to $R_{\mathcal{E}}(0)$)

f-domain expression for the limit model

Combine:

$$\varepsilon(t, \theta) = H(q, \theta)^{-1} [y(t) - G(q, \theta)u(t)]$$

with

$$y(t) = G_0(q)u(t) + v(t)$$

then follows:

$$\varepsilon(t, \theta) = H(q, \theta)^{-1} [[G_0(q) - G(q, \theta)]u(t) + v(t)]$$

Consequence of the identification criterion:

$$\bar{V}(\theta) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\varepsilon}(\omega) d\omega =$$

$$\bar{V}(\theta) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{|G_0(e^{i\omega}) - G(e^{i\omega}, \theta)|^2 \Phi_u(\omega) + \Phi_v(\omega)}{|H(e^{i\omega}, \theta)|^2} d\omega$$

This criterion plays the role of approximation criterion.

Alternative form

Write

$$\varepsilon(t, \theta) = e(t) + \frac{G_0(q) - G(q, \theta)}{H(q, \theta)} u(t) + \frac{H_0(q) - H(q, \theta)}{H(q, \theta)} e(t)$$

then

$$\theta^* = \arg \min_{\theta}$$

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} \left[\frac{|G_0(e^{i\omega}) - G(e^{i\omega}, \theta)|^2}{|H(e^{i\omega}, \theta)|^2} \Phi_u(\omega) + \frac{|H_0(e^{i\omega}) - H(e^{i\omega}, \theta)|^2}{|H(e^{i\omega}, \theta)|^2} \sigma_e^2 \right] d\omega$$

Expression shows how θ^* is obtained.

Two mechanisms:

- ▶ Minimization of $\frac{|G_0(e^{i\omega}) - G(e^{i\omega}, \theta)|^2 \Phi_u(\omega)}{|H(e^{i\omega}, \theta)|^2}$
- ▶ Minimization of $\frac{|H_0(e^{i\omega}) - H(e^{i\omega}, \theta)|^2 \sigma_e^2}{|H(e^{i\omega}, \theta)|^2}$

Problems are coupled if $H(q, \theta)$ is parametrized.

Observation:

Type of approximation of G_0 by $G(q, \hat{\theta}_N)$ is dependent on noise model $H(q, \theta)$.

Special case: fixed noise model $H(q, \theta) = H_*(q)$ (e.g. OE).

Then

$$\theta^* = \arg \min_{\theta} \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{|G_0(e^{i\omega}) - G(e^{i\omega}, \theta)|^2 \Phi_u(\omega)}{|H_*(e^{i\omega})|^2} d\omega$$

(second term in integrand is θ -independent).

θ^* is determined by minimizing the integrated quadratic error $G_0 - G(\theta)$ with weighting function

$$\frac{\Phi_u(\omega)}{|H_*(e^{i\omega})|^2}$$

At those frequencies where the weighting function is large, the model error will be small.

Example

$$y(t) = G_0(q)u(t)$$

G_0 5th order; $\Phi_u(\omega) = 1$ (white noise).

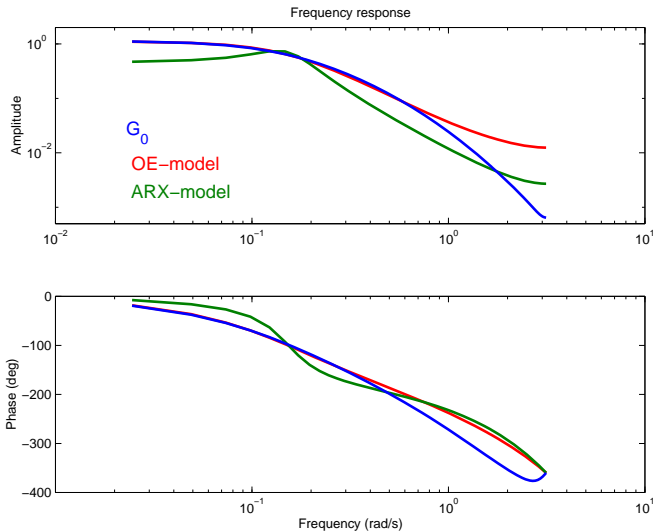
To illustrate the effect of approximation 2nd order models will be estimated:

- ▶ OE-model, 2nd order

$$y(t) = \frac{b_1q^{-1} + b_2q^{-2}}{1 + f_1q^{-1} + f_2q^{-2}}u(t) + e(t)$$

- ▶ ARX-model, 2nd order

$$(1 + a_1q^{-1} + a_2q^{-2})y(t) = (b_1q^{-1} + b_2q^{-2})u(t) + e(t)$$



Comparison of situations:

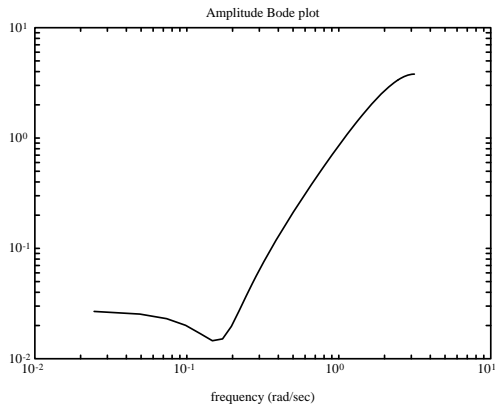
OE-model

$$\min \frac{1}{2\pi} \int_{-\pi}^{\pi} |G_0 - G(\theta)|^2 \Phi_u d\omega$$

ARX-model

$$\min \frac{1}{2\pi} \int_{-\pi}^{\pi} |G_0 - G(\theta)|^2 \Phi_u \cdot |A(e^{i\omega}, \theta)|^2 d\omega$$

Additional weighting with (a priori unknown) function.



Weighting function $|A(e^{i\omega}, \hat{\theta}_N)|$ in ARX-case.

General situation:

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{|G_0(e^{i\omega}) - G(e^{i\omega}, \theta)|^2 \Phi_u(\omega) + \Phi_v(\omega)}{|H(e^{i\omega}, \theta)|^2} d\omega$$

In case of independent parametrization or fixed noise model

$H = H_*$:

$G(q, \hat{\rho}_N)$ is obtained through

$$\hat{\rho}_N = \arg \min \frac{1}{2\pi} \int_{-\pi}^{\pi} |G_0 - G(\rho)|^2 \frac{\Phi_u}{|H_*|^2} d\omega$$

This holds for OE, BJ, FIR.

In other cases: compromise in which a.o. Φ_v is playing a role in the approximation of G_0 .

Prefiltering of data

Prefiltering of the data with filter $L(q)$, leads to

$$\varepsilon_F(t, \theta) = L(q)\varepsilon(t, \theta)$$

and

$$\Phi_{\varepsilon_F}(\omega) = |L(e^{i\omega})|^2 \cdot \Phi_{\varepsilon}(\omega)$$

For a fixed noise model the weighting function becomes:

$$\frac{\Phi_u(\omega)|L(e^{i\omega})|^2}{|H_*(e^{i\omega})|^2}$$

Influencing the approximation criterion
(and so the resulting model)
by

- ▶ Choice of input spectrum Φ_u
- ▶ Choice of prefilter L
- ▶ Choice of noise model H

Example

$$\mathcal{S}: y(t) = G_0(q)u(t) + e(t)$$

with G_0 4th order with three delays.

We have to use a given set of data Z^N ($N = 5000$) for the identification where u is the sum of a white noise of variance 5 and four high-frequency sinusoids with amplitude 10, with frequencies: 1, 1.3, 1.5 and 2.0 rad/sec.

Objective: Using the given data, identify a good model $G(q, \hat{\theta}_N)$ for G_0 in the frequency range $[0 \ 0.7]$ rad/sec in the reduced order model structure:

$$\mathcal{M} = OE(n_b = 2, n_f = 2, n_k = 3)$$

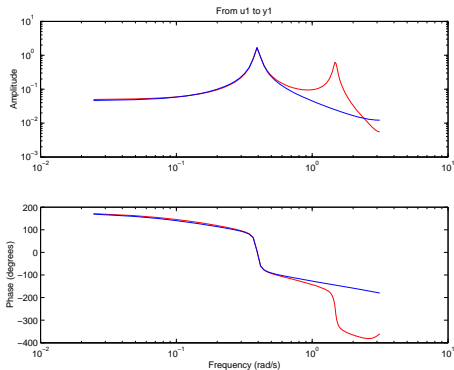
Since Z^N is given, the only degree of freedom we have is to use a pre-filter L to shape the bias error.

We want a small bias error in the frequency range $[0 \ 0.7] \implies$ choose L such that $|L(e^{i\omega})|^2 \Phi_u(\omega)$ is relatively (much) larger in the frequency range $[0 \ 0.7]$ than in $[0.7 \ \pi]$.

$\implies L$ Butterworth low pass filter of order 7 and cut-off frequency 0.7 rad/s

We filter u and y collected from \mathcal{S} by this L and we obtain filtered data with which we perform the identification in \mathcal{M}

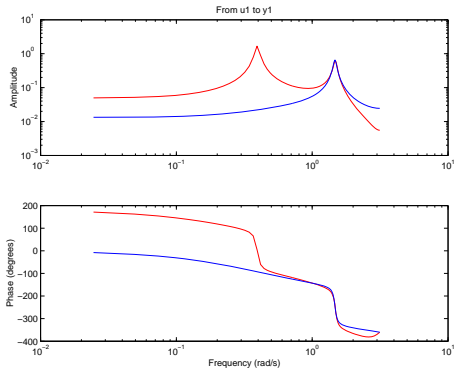
G_0 (red) and $G(\hat{\theta}_N)$ (blue) identified with the filtered data.



$\Rightarrow G(\hat{\theta}_N)$ is OK.

What if we do not use a pre-filter L ?

G_0 (red) and $G(\hat{\theta}_N)$ (blue) identified with the data in Z^N



$\Rightarrow G(\hat{\theta}_N)$ is KO

Summary

- The linearity-in-the-parameters property of ARX and FIR models ([linear regression](#)), leads to simple (convex) computational identification algorithms with an [analytical solution](#).
- Consistency result for these model structures follow the earlier presented consistency results.
- If model sets can not capture reality, i.e. $\mathcal{S} \notin \mathcal{M}$, or $G_0 \notin \mathcal{G}$, the asymptotically estimated models can be characterized with an [approximation criterion](#).
- [Approximative properties of identified models](#) can be influenced by [input signal](#), [prefilter](#) and [noise model](#).